

**ORIGINAL ARTICLE****The Relationship between Resettlement and Birth Rates: The Case of Gambella, Ethiopia****Aynalem Adugna<sup>1</sup>, Helmut Kloos<sup>2</sup>****ABSTRACT**

**BACKGROUND:** This study aims to examine the possible impacts of resettlement on birth rates by using the length of stay variable in the 2000 Demographic and Health Survey (DHS). **METHODS:** Data in all three rounds of Gambella Administrative Region's Demographic and Health Surveys (DHS) are analyzed. The neighboring administrative region of Benishangul-Gumuz is used as a control. The multivariate analysis of variance (MANOVA) is applied with duration of residence as a categorical independent variable. The statistical software SAS is used. **RESULTS:** In a univariate analysis of Gambella's DHS 2000, duration of residence has a significant effect on mothers' age at first birth ( $p < 0.001$ ), the number of children born within the five years of the survey ( $p < 0.001$ ), and the total number of children ever born ( $P < 0.001$ ). In the MANOVA analysis, the duration effect on all three is also statistically significant ( $p < 0.001$ ). **DISCUSSION:** Resettlement had a disruptive effect on birth rates among females who were just coming into marriageable ages in places of origin but were resettled to Gambella. Although the disruptive effects waned over time, the initial shortfall resulted in reduced overall lifetime births for settler women who were not past the midpoint of their reproductive years at arrival. **CONCLUSION:** Based on the reproductive history of female settlers with different duration of residence in the resettlement schemes, we recommend the reinstatement of the length of residence question in future DHS surveys in Ethiopia to allow a longitudinal tracking of demographic trends among nonnative populations.

DOI: <http://dx.doi.org/10.4314/ejhs.v26i4.8>**INTRODUCTION**

Three Demographic and Health Surveys (DHS) were conducted in Ethiopia, in 2000, 2005 and 2011. Results of these surveys were reported at the national level (1-3), allowing examinations of demographic patterns (4-9) and topic-specific analyses (10,11). This has helped close some knowledge gaps. However, the demographic impacts of population mobility on fertility and mortality have not been analyzed. As a result, the short- and long-term demographic impacts of migrations, in particular, the impacts of government-sponsored resettlement of the 1980s, 1990s, and 2000s remain largely unknown. This study aims to help fill this knowledge gap using a point-in-time analysis together with a longitudinal

approach that compares reported numbers of births for settlers and non-settlers in Gambella Administrative Region. Gambella Region was selected because it tops the list of destinations of government-sponsored resettlement dating back to the 1970s (12). The neighboring region of Benishangul-Gumuz served as a control (Figure 1). Both regions have hosted settlers, refugees and internally displaced persons over the years, and have similar lowland climates and vegetations (13-20), but Gambella's proportion of the settler population is much larger.

The study has two main objectives: 1) to examine the possible impacts of government-sponsored resettlement on birthrates in Gambella Administrative Region, and 2) to show the

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usefulness of using a neighboring administrative region with similar demography, socio-economic and environmental characteristics as control (Figure 1).



*Figure 1: Map showing the study and control areas*

Table 1: Number and percentage of women respondents in Gambella, Benishangul- Gumuz, and nationally by 5 year age group (Ethiopian Demographic Survey, 2011)

5-year age groups	Nation		Benishangul-Gumuz		Gambella	
	Number	Percent	Number	Percent	Number	Percent
15 - 19	3,835	23.22	554	24.83	166	20.15
20 - 24	3,022	18.30	412	18.47	145	17.60
25 - 29	3,185	19.29	429	19.23	165	20.02
30 - 34	2,100	12.72	283	12.68	132	16.02
35 -39	1,958	11.86	241	10.80	98	11.89
40 - 44	1,314	7.96	152	6.81	65	7.89
45 - 49	1,101	6.67	160	7.17	53	6.43
<b>Total</b>	<b>16,515</b>	<b>100.0</b>	<b>2,231</b>	<b>100.0</b>	<b>824</b>	<b>100.0</b>

## MATERIALS AND METHODS

We used a longitudinal approach and a point-in-time multivariate analysis in which women were placed in five-year age cohorts. The sample-based DHS 2000, 2005, and 2011 data were obtained from [http://dhsprogram.com/data/dataset/admin/login\\_main.cfm](http://dhsprogram.com/data/dataset/admin/login_main.cfm) following an online application for download. The enumeration areas in Ethiopia's last two censuses served as the sampling frame (1-3). Complete SAS and GIS/GPS data sets were downloaded. There were 876 women between the age of 15-49 in Gambella's 2000 sample, 921 women in the 2005 sample, and 1,215 women in the 2011 sample. Three variables were analyzed: 1) age of women at the birth of the first child, 2) the number of

children born within five years of the 2000 and 2005 surveys, and 3) the number of children ever born. The 2000 and 2005 DHS have duration of residence information but the 2011 DHS does not. Six records with missing duration of residence values were excluded in the 2000 sample, and 11 cases were excluded in the 2005 sample. Comparable DHS data were used for the control region Benishangul-Gumuz.

The statistical software SAS was used to conduct a multivariate analysis of variance (MANOVA). MANOVA is simply an analysis of variance (ANOVA) with two or more dependent variables. Whereas ANOVA tests for the difference in means between two or more groups, MANOVA tests for the difference in two or more

vectors of means (21). Wilk's Lambda is one of the measures. It can be interpreted as the proportion of the variance in the outcomes that is not explained by an effect (21). To calculate Wilk's Lambda for each characteristic root, one has to calculate  $1/(1 + \text{the characteristic root})$ , then find the product of these ratios. Pillai's Trace is the other measure (21). To calculate Pillai's trace, each characteristic root is divided by  $1 + \text{the characteristic root}$ , and then the sums of these ratios are summarized. Hotelling-Lawley Trace is very similar to Pillai's Trace. It is the sum of the roots of the product of the sum-of-squares matrix of the model and the sum-of-squares matrix of the error for the two linear regression functions and is a direct generalization of the F statistic in ANOVA (21). The Hotelling-Lawley Trace is calculated by summing the characteristic roots listed in the output. Roy's Greatest Root is the largest of the roots of the product of the sum-of-squares matrix of the model and the sum-of-squares matrix of the error for the two linear regression functions. The value of Roy's Greatest Root is the largest of the characteristic roots. Because it is a maximum, it can behave differently from the other three test statistics. As a general rule, in instances where the other three are not significant and Roy's is significant, the effect should be considered non-significant (Table 1).

## RESULTS

**Age distribution of sample women:** Figure 1 shows the percentage distribution of sample women in the 2011 DHS by five-year age groups. Differences in age distribution between women interviewees nationally, those in Gambella and those in Benishangul-Gumuz were not statistically significant (Figure 1). The proportion of women under 30 years of age was 60.8% nationally, 62.5% in Benishangul-Gumuz, and 57.8% in Gambella. The main difference is in the length of time various age cohorts of women had lived in their places of interview by the time of DHS 2011. Analysis of data from all three surveys showed significantly higher proportions of nonnatives in Gambella's 2011 DHS than nationally or in Benishangul-Gumuz.

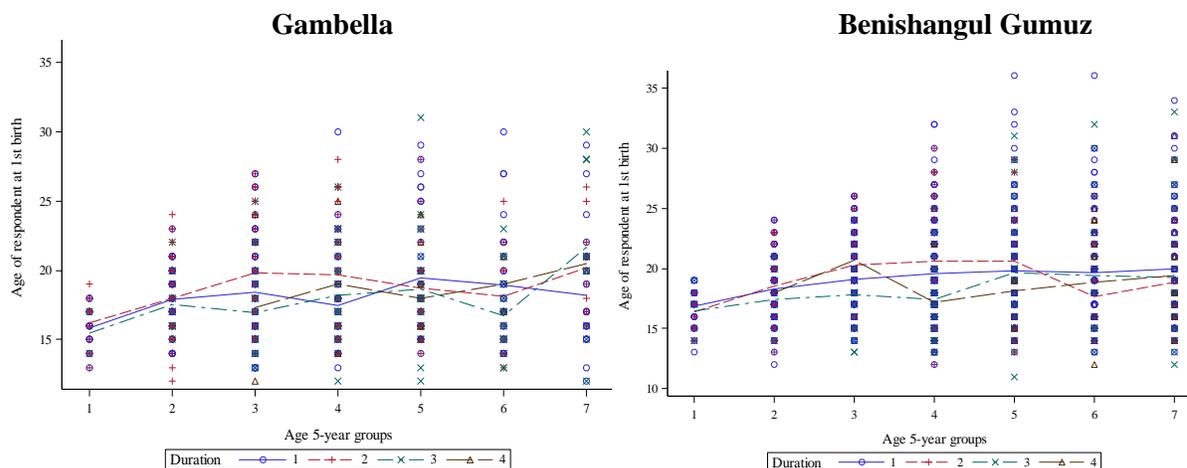
**Duration of residence:** Results of DHS 2000 show that 15-19 year old nonnatives represented 20.1% of the population nation-wide, 19.1% in Benishangul-Gumuz, and 41.8% in Gambella. For 20-24 year olds, the proportions of nonnatives at the national level, in Benishangul-Gumuz, and in Gambella were 35.3, 29.6 and 66.6% respectively, suggesting a minority status for 20-24 year old females in Gambella but not nationally or in Benishangul-Gumuz. For the 25-29 year old females, the respective proportions of nonnative females were 42.7, 39.0, and 65.6%. Nonnative 30-34 year old females were also a majority in Gambella (54.4%) but a minority nationally (43.7%) and in Benishangul-Gumuz (45.6%).

By the time of DHS 2000, the proportion of females residing in places of interview for less than 10 years was 28.0% nationally, 23.2% in Benishangul-Gumuz, and 44.0% in Gambella. The proportions with length of stay between 10 and 19 years were 7.0% nationally, 5.6% in Benishangul-Gumuz, and 12.6% in Gambella.

**Duration of residence and age of respondents at the birth of their first child:** The dots in Figure 2 show how old an individual female interviewee was at the birth of her first child. The line graphs show the average age of women at the birth of the first child by duration of residence. A focus on the extreme age groups (15-19 and 45-49) reveals generational differences in the effect of duration on age at first birth. Three observations emerge: a) in both Gambella and Benishangul-Gumuz, there is no difference in age at first birth between natives and nonnatives for the 15-19 age group, b) in both Gambella and Benishangul-Gumuz, there is a difference in age at first birth between natives and nonnatives for the 45-49 age group and c) the difference is in opposite directions with native 45-49 year olds having a lower average age at first birth than nonnatives in any duration group in Gambella while nonnatives had a higher age at first birth than natives in any duration group in Benishangul-Gumuz. The effect of duration of residence on age at first birth in Gambella is statistically significant ( $p < .0001$ ) (Table 1).

**Legend**

Age group            1 : 15-19            2 : 20-24            3 : 25-29            4 : 30-34  
 Duration (years)   1: Native            2: 0-9                3: 10-19            4: 20+



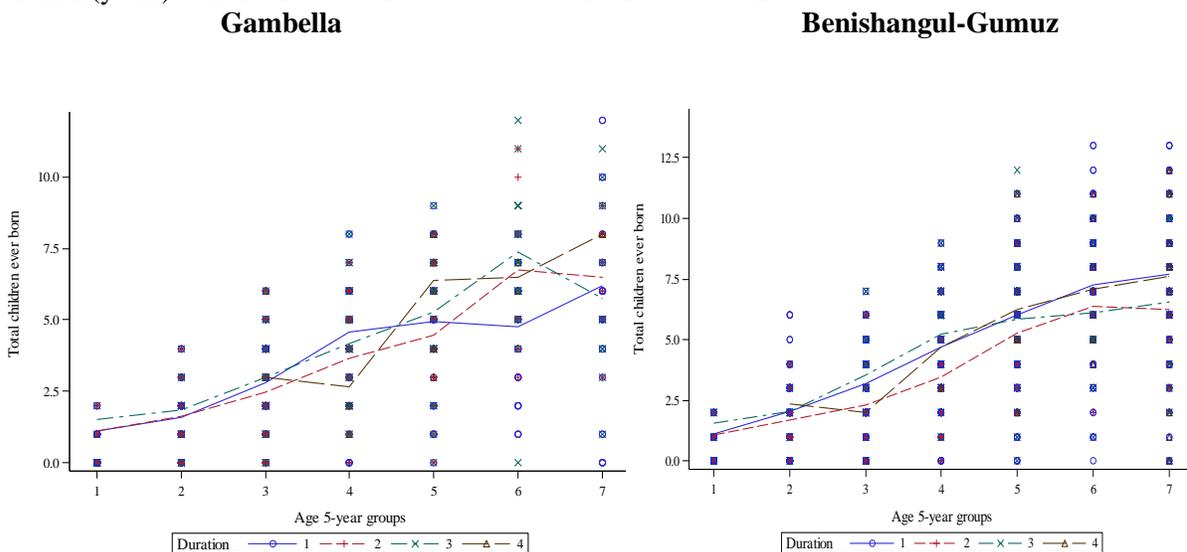
**Figure 2:** Age of respondent at first birth by age at the time of survey and duration of residence (DHS, 2000)

**Duration of residence and childbearing during the five years prior to DHS 2000:** One fifth (20.7%) of respondents had two or more births during the five years prior to DHS 2000 nationally, 24.9% in Benishangul-Gumuz and 15.6% in Gambella. On the other hand, the proportion of female respondents who only had one birth during

that period was 26.5% nationally, 26.7% in Benishangul-Gumuz, and 36.6% in Gambella. For Gambella, the difference by duration of residence is statistically significant ( $p < .0001$ ) (Table 1). Figure 3 reveals significant effect of duration of residence in Gambella, but not in Benishangul-Gumuz.

**Legend**

Age group            1 : 15-19            2 : 20-24            3 : 25-29            4 : 30-34  
 Duration (years)   1: Native            2: 0-9                3: 10-19            4: 20+



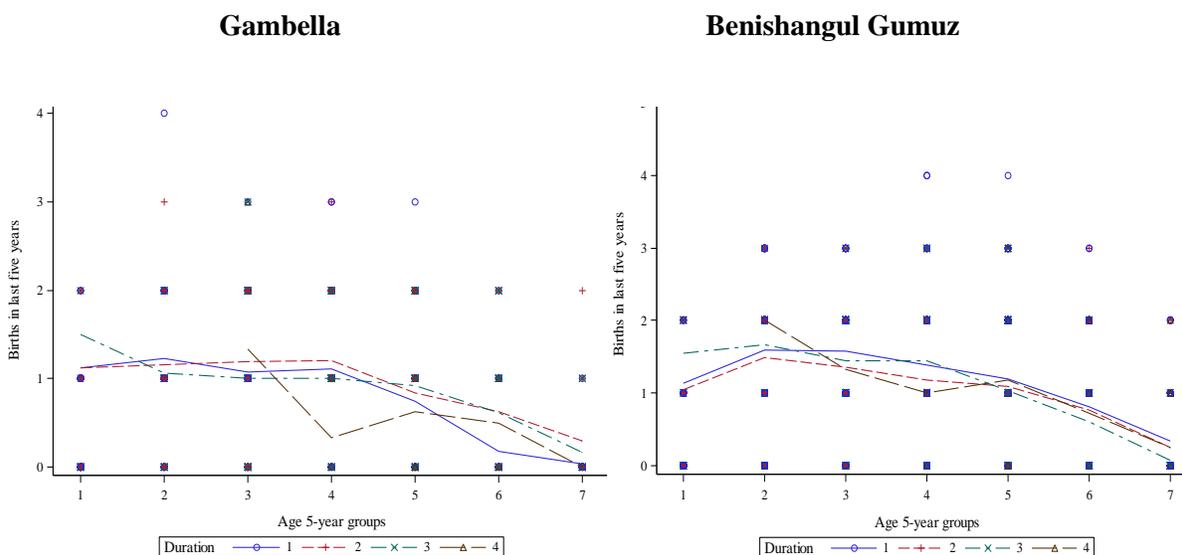
**Figure 3:** Total children ever born by age at the time of survey and duration of residence

Nonnatives in the 30-34 age group were between the ages of 10 and 14 at the time of arrival prior to 1980. The 2.0 children per woman deficit for the 30-34 year-olds at DHS 2000 could be explained by the disruptive effects of resettlement. There is no native-nonnative difference for the two age groups in Benishangul-Gumuz. The native-nonnative difference in Gambella also applies to

the 35-39, 40-44 and 45-49 year age groups, with nonnatives reporting approximately two additional lifetime births than natives. This pattern is largely absent in Benishangul-Gumuz, where most curves representing the average number of children by women's age and duration cohorts are overlapping (Figure 4).

**Legend**

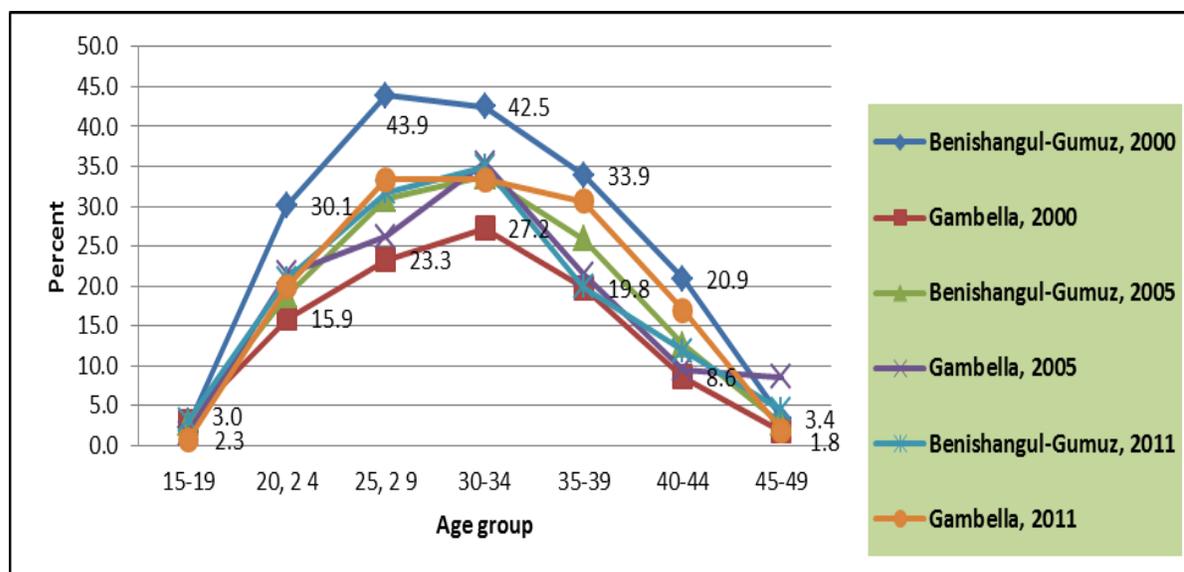
Age group	1 : 15-19	2 : 20-24	3 : 25-29	4 : 30-34
Duration (years)	1: Native	2: 0-9	3: 10-19	4: 20+



**Figure 4:** Number of births in the five years prior to DHS 2000 by age and by duration of residence

Figure 5 presents a comprehensive longitudinal picture. The table portion combines findings from all the three demographic surveys by using the number of births in the last five years variable for the entire 16-year period (the five years preceding DHS 2000 to the end of DHS 2011), and by tracking age cohorts of women. The six-year-interval for 2005 to 2011 was made possible by reclassification of single age distributions using SAS. This approach provided an opportunity to follow recent reproductive histories of an age cohort using the number of births in the last five

years variable as proxy. The graph portion of Figure 5 shows the percentages of women who had two or more births in the five years prior to DHS 2000, DHS 2005, and DHS 2011. Gambella's position during the three surveys is noteworthy for the shift from a bottom position in 2000 to the second highest in 2011. For instance, 43.9% and 23.3% of women aged 25-29 in Benishangul-Gumuz and Gambella had two or more births respectively during DHS 2000. At 33% each, the proportions for the two regions had become equal by DHS 2011.



**Figure 5:** Percentage of women who had two or more births in the five years prior to DHS 2000, DHS 2005, and DHS 2011 (top graph) and the number of children born per woman by age group of respondents between 1995 and 2011 (bottom table)

Table 2: Number of births per woman in DHS 2000, 2005, 2011 by age group of respondents

<b>Benishangul-Gumuz</b>			
Age_15 -19, 2000	<b>0.13</b>		<b>16 Year total</b>
	Age 20-24, 2005	<b>0.50</b>	
		Age 26-30, 2011	<b>0.69</b>
Age 20 -24, 2000	<b>0.56</b>		<b>1.31</b>
	Age 25-29, 2005	<b>0.68</b>	
		Age 31-35, 2011	<b>0.71</b>
Age 25 -29, 2000	<b>0.75</b>		<b>1.95</b>
	Age 30-34, 2005	<b>0.71</b>	
		Age 36-40, 2011	<b>0.57</b>
Age 30 -34, 2000	<b>0.82</b>		<b>2.04</b>
	Age 35-39, 2005	<b>0.59</b>	
		Age 41-45, 2011	<b>0.38</b>
			<b>1.78</b>
<b>Gambella</b>			
Age_15 -19, 2000	<b>0.21</b>		<b>16 Year total</b>
	Age 20-24, 2005	<b>0.46</b>	
		Age 26-30, 2011	<b>0.73</b>
Age 20 -24, 2000	<b>0.66</b>		<b>1.41</b>
	Age 25-29, 2005	<b>0.60</b>	
		Age 31-35, 2011	<b>0.77</b>
Age 25 -29, 2000	<b>0.74</b>		<b>2.02</b>
	Age 30-34, 2005	<b>0.65</b>	
		Age 36-40, 2011	<b>0.68</b>
Age 30 -34, 2000	<b>0.72</b>		<b>2.07</b>
	Age 35-39, 2005	<b>0.57</b>	
		Age 41-45, 2011	<b>0.51</b>
			<b>1.80</b>

**Univariate and multivariate analyses of variance:** A summary univariate analysis based on a point-in-time DHS 2000 data is provided for Gambella followed by a multivariate analysis of variance (MANOVA) (Table 1). Results of the MANOVA with all the three dependent variables in combination and the duration of residence as the independent variable show a statistically significant effect. MANOVA calculates four multivariate test statistics based on the characteristic roots (21). The null hypothesis for each of these tests is the same: the independent variable has no effect on any of the dependent variables. Duration of residence (coded V013) is the independent variable in both the univariate and the multivariate analysis. The first dependent

variable considered was total number of children ever born (code V201). We obtained an R-Square of 0.53, a Root MSE of 1.65, a coefficient of variation of 45.20 and a mean value of 3.65. The second dependent variable in the univariate analysis was number of births in the five years prior to DHS 2000 (code V208), giving a R-Square of 0.24, a Root MSE of 0.67, a coefficient of variation of 72.48 and a mean of 0.92. The third dependent variable in the univariate analysis was age of respondent at first birth (code V212) (R-Square 0.11, Root MSE 3.47, coefficient of variation 18.88, Mean of 18.42). All three results were statistically significant ( $p < 0.0001$ ) (Table 1).

Table 3. Analysis of Variance with age at first birth, children born within 5 years of DHS 2000, and the number of children ever born as dependent variables and duration of residence as the independent variable: Gambella DHS 2000  
 Analysis of Variance: One Dependent and one Independent Variable  
 Independent Variables = Duration of residence (V013)

Dependent Variable: V201 Total children ever born					
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	29	1910.059229	65.864111	24.17	<.0001
Error	628	1711.243202	2.724910		
<b>Corrected Total</b>	<b>657</b>	<b>3621.302432</b>			
Dependent Variable: V208 Births in last five years					
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	29	88.0252326	3.0353528	6.81	<.0001
Error	628	279.8653449	0.4456455		
<b>Corrected Total</b>	<b>657</b>	<b>367.8905775</b>			
Dependent Variable: V212 Age of respondent at 1st birth					
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	29	939.402796	32.393200	2.68	<.0001
Error	628	7592.665593	12.090232		
<b>Corrected Total</b>	<b>657</b>	<b>8532.068389</b>			
MANOVA Test Criteria and F Approximations for the Hypothesis of No Overall Duration (V013) Effect					
Characteristic Root	Percent	Characteristic Vector V'EV=1			
		V201	V208	V212	
<b>3.15696719</b>	<b>93.38</b>	0.03111052	-0.04765103	0.00897116	
<b>0.12704330</b>	<b>3.76</b>	0.00507320	0.03986532	0.00811006	
<b>0.09667761</b>	<b>2.86</b>	0.00703988	0.02893753	-0.00740575	
<b>S=3 M=12.5 N=312</b>					
Statistic	Value	F Value	Num DF	Den DF	Pr > F
<b>Wilks' Lambda</b>	0.19462733	15.68	87	1873.9	<.0001
<b>Pillai's Trace</b>	0.96031765	10.20	87	1884	<.0001
<b>Hotelling-Lawley Trace</b>	3.38068810	24.28	87	1677	<.0001
<b>Roy's Greatest Root</b>	3.15696719	68.36	29	628	<.0001

**NOTE: F Statistic for Roy's Greatest Root is an upper bound.**

## DISCUSSION

No studies are available regarding the impacts of large scale resettlement on birth rates. By achieving its two main objectives of showing the fertility effects of resettlement in Gambella and of demonstrating the advantages of using a neighboring region as control, this study can help fill the knowledge gap. We adopted a dual approach of longitudinal age-cohort tracking on the one hand and a point-in-time multivariate analysis on the other, to examine the possible demographic impacts of resettlement in Gambella Region. Both the dual approach and the use of Benishangul-Gumuz as control represent a methodological improvement in the analysis of DHS data. Our overall finding is that resettlement has the effect of reducing the number of lifetime births for women who arrived at young ages, but not on women who were past the mid-point of their reproductive years at the time of arrival. The effect of resettlement on women just coming into marriageable ages at the time of arrival is unlikely to be the singular reason Gambella has a lower total fertility rate (TFR) than Benishangul-Gumuz (1-3). We are not able to show whether or not settler-native differentials existed in contraceptive prevalence (5), or if Gambella's sociopolitical factors (13, 15, 16 and 18) differentially affected birth rates among settlers.

We analyzed reproductive histories using a count of lifetime births and recent fertility based on the 16-year period starting with five years prior to DHS 2000 and ending at the conclusion of DHS 2011. The picture for Gambella is complex, and specific results depended on whether the focus was on cumulative numbers of lifetime births or on recent fertility. Our analysis traced the reproductive histories of the 1970s and 80s arrivals based on age at the time of DHS 2000. The comparison between women in the 30-34 and 40-44 age groups is particularly noteworthy. For 40-44 year-olds, the total number of births is higher among nonnatives in the 20+ duration group than among natives by about 1.5 lifetime births per woman. The reverse is true for 30-34 year olds in the 20+ years' duration group. Here, natives have about 2.0 additional lifetime births than nonnatives. This is not surprising given that nonnatives in the 30-34 age groups were aged 10-14 at the time of arrival prior to 1980. The 2.0

lifetime births per woman deficit for the 30-34 year old nonnatives could be thought of as resulting from the disruptive effects of resettlement. Supporting evidence comes from the absence of native-nonnative differences for the two age groups in our control region, Benishangul-Gumuz. The higher cumulative lifetime births in Gambella also apply to 35-39, 40-44 and 45-49 year-olds where approximately two additional births were registered for nonnatives than for natives. This difference is largely absent in Benishangul-Gumuz. In other words, the disruptive effect of resettlement on child birth was limited to women who were younger than 35 years of age at DHS 2000. Settler women aged 40 or higher at DHS 2000 had higher lifetime births in any duration group than native born women. These findings do not apply to Benishangul-Gumuz where native-born women have higher lifetime births than nonnatives in any age or duration group.

A different portrait emerges when recent reproductive performances (births in the five years preceding the three surveys) are analyzed. The picture from Gambella's DHS 2000 is one of an initial deficit for age cohorts of nonnatives compared to natives; a deficit which is then made up over a period of 16 years (five years prior to DHS 2000 through the end of DHS 2011).

Results of the multivariate analysis with duration of residence as the independent variable and three dependent variables (age at the time of first birth, number of births in the five years within five years of the 2000 and 2005 surveys and lifetime births) showed a statistically significant effect of duration of residence when based on a point-in-time analysis of Gambella's DHS 2000. However, a longitudinal age cohort analysis showed the duration effect as moderating over time. The moderation is reflected in the absence of significant native-nonnative differences in the number of children ever born when five-year age cohorts are followed from DHS 2000 to the DHS 2011.

The longitudinal approach followed in Figure 5 shows the success settler women in Gambella achieved in making up for the deficits recorded in DHS 2000. The right most columns for Gambella and Benishangul-Gumuz reveal this. For example, 20-24 year old females in Benishangul-Gumuz started with a five-year

average of 0.66 births in 2000 and ended with a cumulated total of 2.02 births for the 16 year period (five years prior to DHS 2000 through DHS 2011). Their counterparts in Gambella started with a lower five year average of 0.56 in 2000 but had matched the 16-year cumulated total for women in Benishangul-Gumuz by the end of the 16-year period in 2011.

We recommend that the length of stay variable be reinstated in Ethiopia's future DHS surveys to capture the effects of population mobility on birth rates and other demographic variables. Monitoring of the possible impacts of mobility on health and death is also dependent on the inclusion of the length of stay variable.

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